

# CELL AVERAGING CFAR DETECTOR WITH SCALE FACTOR CORRECTION THROUGH THE METHOD OF MOMENTS FOR THE LOG-NORMAL DISTRIBUTION

## Detector CFAR de Promediación con Corrección del Factor de Ajuste a través del Método de los Momentos para la Distribución Log-Normal

**José Raúl Machado Fernández**

Ingeniero en Telecomunicaciones y Electrónica, Master en Telecomunicaciones y Electrónica, Doctorante, Profesor e Investigador, Grupo de Investigación de Radares, Departamento de Telecomunicaciones y Telemática, Facultad de Eléctrica, Instituto Superior Politécnico José Antonio Echeverría (ISPJAE-CUJAE), La Habana, Cuba,  
[josemf@electronica.cujae.edu.cu](mailto:josemf@electronica.cujae.edu.cu)

**Jesús C. Bacallao Vidal**

Ingeniero Eléctrico, Doctor en Ciencias Técnicas, Profesor Titular e Investigador, 2do Jefe del Grupo de Investigación de Radares, Departamento de Telecomunicaciones y Telemática, Facultad de Eléctrica, Instituto Superior Politécnico José Antonio Echeverría (ISPJAE-CUJAE), La Habana, Cuba,  
[bacallao@electronica.cujae.edu.cu](mailto:bacallao@electronica.cujae.edu.cu)

### ABSTRACT

The new LN-MoM-CA-CFAR detector is introduced, exhibiting a reduced deviation of the operational false alarm probability from the value conceived in the design. The solution solves a fundamental problem of CFAR processors that has been ignored by most of its predecessors. Indeed, most of the previously proposed schemes deal with sudden changes in the clutter level, whereas the new solution has an improved performance against statistical slow changes that occur in the background signal. It has been proved that these slow changes have a remarkable influence on the selection of the CFAR adjustment factor, and consequently in maintaining the false alarm probability. The authors took advantage of the high precision achieved by the MoM (Method of Moments) in the estimation of the Log-Normal (LN) shape parameter, and the wide application of this distribution to radar clutter modelling, to create an architecture that offers precise results and it's computationally inexpensive at the same time. After an intensive processing, involving 100 million Log-Normal samples, a scheme, which operates with excellent stability reaching a deviation of only 0,2884% for the probability of false alarm of 0,01, was created, improving the classical CA-CFAR detector through the continuous correction of its scale factor.

**Keywords:** Method of Moments, CFAR Detectors, False Alarm Probability, Log-Normal Distribution, Radar Clutter

## RESUMEN

Se presenta el nuevo detector LN-MoM-CA-CFAR que tiene una desviación reducida en la tasa de probabilidad de falsa alarma operacional con respecto al valor concebido de diseño. La solución corrige un problema fundamental de los procesadores CFAR que ha sido ignorado en múltiples desarrollos. En efecto, la mayoría de los esquemas previamente propuestos tratan con los cambios bruscos del nivel del *clutter* mientras que la presente solución corrige los cambios lentos estadísticos de la señal de fondo. Se ha demostrado que estos tienen una influencia marcada en la selección del factor de ajuste multiplicativo CFAR, y consecuentemente en el mantenimiento de la probabilidad de falsa alarma. Los autores aprovecharon la alta precisión que se alcanza en la estimación del parámetro de forma Log-Normal con el MoM, y la amplia aplicación de esta distribución en la modelación del *clutter*, para crear una arquitectura que ofrece resultados precisos y con bajo costo computacional. Luego de un procesamiento intensivo de 100 millones de muestras Log-Normal, se creó un esquema que, mejorando el desempeño del clásico CA-CFAR a través de la corrección continua de su factor de ajuste, opera con una excelente estabilidad alcanzando una desviación de solamente 0,2884% para la probabilidad de falsa alarma de 0,01.

**Palabras Claves:** Método de los Momentos, Detectores CFAR, Probabilidad de Falsa Alarma, Distribución Log-Normal, Clutter de Radar.

## INTRODUCTION

A radar is a device that emits electromagnetic waves and gather the resulting echo caused by objectives in the proximity [1]. The mission of the radar is to detect nearby targets of interest and to discard those that do not relate with the particular application for which it's designed. Therefore, some objectives (such as clouds) can be considered as targets for certain applications (meteorology) and as interfering signal (military) for others [2].

Echo signals produced at ground surface, sea surface or weather volumes (clouds, fog, rain) are assumed as interference and called clutter, in most radar applications [3]. The magnitude of the clutter signal cannot be deducted by purely deterministic mechanisms. Hence, its modeling falls in the field of statistics.

The false alarm probability ( $P_f$ ) is one of the radar fundamental parameters because it defines the frequency of occurrence of type 1 errors. A type I error takes place when a clutter signal is mistakenly classified as a target [4].

Although decreasing the  $P_f$  is beneficial, generally it has the undesirable effect of disturbing the probability of detection ( $P_d$ ), which is another essential parameter. To resolve this 'balance' relation, the Neyman-Pearson criteria is applied, which states that a certain level of  $P_f$  must be guaranteed first as an essential requirement. Only afterwards, actions can be taken to improve the  $P_d$

by modifying. Therefore, radar detectors or processors must have the CFAR (Constant False Alarm Rate) property, because they must ensure a constant  $P_f$  is maintained the entire period of operation [5].

The classical radar detector, known as CA-CFAR (Cell Averaging CFAR), uses a sliding window with multiple slots that it's passed through the whole coverage area. Thus, each resolution cell has a chance to occupy the central position of the window. The processor evaluates the cell in this position deciding on the presence or absence of a target. The information in the neighboring cells is used to calculate the average of the clutter, which is then multiplied by a scale or adjustment factor ( $T$ ), resulting in the establishment of a detection threshold. If the magnitude of the sample in the central slot is greater than the threshold, a target is considered to be present [6].

## **MOTIVATION AND OBJECTIVE**

While the CA-CFAR guarantees a constant  $P_f$  for several radar scenarios, the occurrence of multiple close targets, and the appearance of sudden changes in the clutter level, alter the mechanism of average calculation, provoking deviations from the design  $P_f$  [6]. To prevent malfunctions against these critical situations, multiple alternative detectors have been created, being all variations from the original CA-CFAR [7-10].

Nevertheless, most authors ignore the effect of clutter slow statistical changes in the maintaining of the  $P_f$ . Their designs lack of an adaptive correction of the multiplicative  $T$  factor and focus exclusively on changing the mechanism for calculating the average.

In contradiction, a lot of papers have verified the need to modify the shape parameter of the distributions used as clutter models for achieving a quality fit with empirical data [11-14]. The variation in the distribution need to be translated into a correction of the CFAR  $T$  factor for guaranteeing the stability of the system. Indeed, in [15] it was demonstrated, by performing simulations with computer-generated data, that a processor with a constant  $T$  value was unable to maintain its design  $P_f$  when it was submitted to clutter whose shape parameter changed over time.

In order to solve the above described problem, the authors aimed to create an improved CFAR detector, able to correct its  $T$  factor according to the clutter statistical variations. To do this, they selected the Log-Normal (LN) distribution as a model for clutter and the CA-CFAR detector as a reference implementation where the improved solution will be introduced. The Method of Moments (MoM) was the mathematical tool used in the estimation of the LN shape parameter. The new scheme was baptized as LN-MoM-CA-CFAR processor.

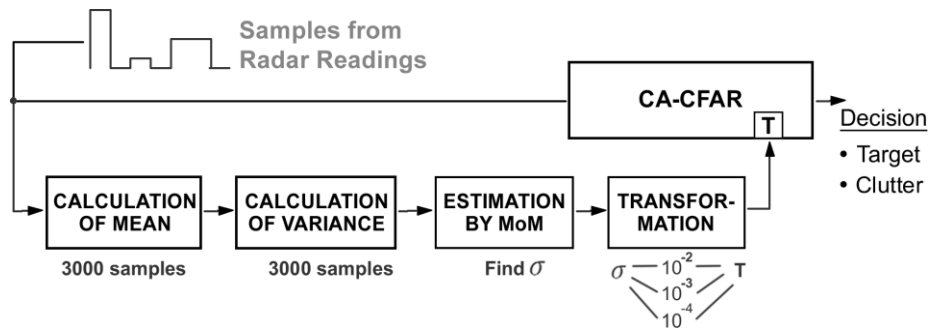
## **1. MATERIALS AND METHODS**

The Materials and Methods section presents the structure of the LN-MoM-CA-CFAR scheme and explains its internal blocks. Special attention is given to the

description of the LN statistical model, the derivation of MoM expressions and the selection of the range of possible values for the shape parameter. Subsequently, the experiments required to test the scheme in MATLAB are described.

### 1.1. THE LN-MOM-CA-CFAR SCHEME

The LN-MoM-CA-CFAR scheme was constructed by adding four blocks to the classical CA-CFAR as shown in Figure 1. Samples of radar readings are fed simultaneously to the input of the CA-CFAR and to the block for calculating the mean. Afterwards, the variance is calculated using, as it was also done for the mean, 3000 statistically independent samples. This amount of samples was selected after consulting procedures followed by other authors [16-18].



**Fig. 1.** Structure of the LN-MoM-CA-CFAR scheme.

**Source:** The authors.

The main advantage of using the LN distribution as preferential clutter model is the straightforward derivation of the parameters by the Method of Moments. Most distributions related to radar clutter (such as Weibull or K) allow a quick estimation by the MoM, but the accuracy of the result is limited. However, the MoM concurs with the expression of ML (Maximum Likelihood) for the LN distribution, thanks to its close relationship with the Normal or Gaussian distribution. So, the estimation of the LN parameters can be made in a quickly and accurately way [19].

In Figure 1, the block "Calculation of Mean" actually estimated the mean of the natural logarithm of the samples. The same goes for the block "Calculation of Variance". These are the requirements for LN MoM formula which is given below.

Once estimated LN shape parameter ( $\sigma$ ), it's necessary to translate it into a viable correction of the  $T$  value. The selection of  $T$  depends on the wanted false alarm rate. Therefore, the authors selected three classic  $P_f$  values to design the solution:  $P_f = 10^{-2}$ ,  $P_f = 10^{-3}$  and  $P_f = 10^{-4}$ . These figures are a common choice in radar related investigations, and were previously used in [20, 21]. The designer may choose any of these  $P_f$ s according to its own requirements.

## 1.2. THE LOG-NORMAL STATISTICAL DISTRIBUTION AND THE MOM

The LN statistical distribution has had an extensive application in radars issues. It's widely used for both sea [12, 13, 16, 22] and ground clutter [23-25]. Even if the model prevails sometimes for a complete set of low grazing angle measurements [16, 26], its best fit generally appears for sub-sets with HH polarization [27], when measuring the data spatial distribution [25], and for cells containing strong targets reflections [16, 28].

In [29, 30] the following expression was used for the LN PDF (Probability Density Function):

$$f(x|\mu, \sigma) = \frac{1}{\sigma x \sqrt{2\pi}} \exp \left[ -\frac{1}{2} \left( \frac{\ln(x) - \mu}{\sigma} \right)^2 \right] \quad (1)$$

Where  $\mu$  and  $\sigma$  are the scale and shape parameters respectively, and  $x$  represents the clutter amplitude measurements. The CDF (Cumulative Density Function) [31] and the moment generating function [31] are given below:

$$F(x) = \frac{1}{2} \left( 1 + \operatorname{erf} \left( \frac{\ln x - \mu_N}{\sigma_N \sqrt{2}} \right) \right) \quad (2)$$

$$M_n = \exp \left( n\mu_N + \frac{(n\sigma_N)^2}{2} \right) \quad (3)$$

Where  $\operatorname{erf}(\cdot)$  is the error function,  $\mu$  is the average amplitude and  $\sigma^2$  is the variance of  $\ln x^2$ .

The MoM derivation for the LN case, addressed in [19], allows to estimate the parameters of the distribution with the following expressions:

$$\mu = \frac{1}{N} \sum_{i=1}^N \ln(x_i) \quad (4)$$

$$\sigma^2 = \frac{1}{N-1} \left\{ \sum_{i=1}^N \ln(x_i)^2 - \mu^2 \right\} \quad (5)$$

Where  $N$  is the total amount of samples,  $i$  is the number of the current sample which goes from 1 to  $N$ , and  $x_i$  is the  $i$ -th sample of the set.

While the LN distribution allows any configuration of its parameters from a mathematical point of view, not all combinations reflect real situations. After a review of the related literature, the authors decided that the  $\sigma$  interval between 0,025 and 1,25 covers most of the operating conditions [24, 27, 32, 33].

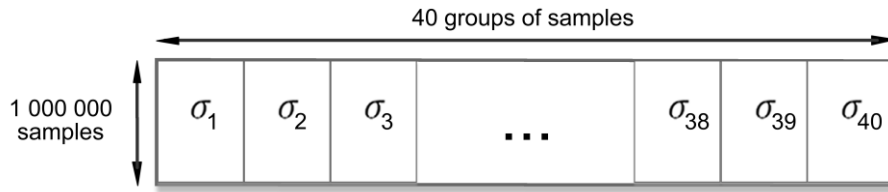
### 1.3. DESIGN OF THE EXPERIMENTS

The mathematical foundations of the proposed scheme were given previously, together with the main precedents of the literature. The following experiments were carried out by the authors to build and test the LN-MoM-CA-CFAR processor.

#### 1.3.1. BUILDING THE TRANSFORMATION BLOCK

Firstly, the authors implemented the scheme from Figure 1 in MATLAB. The main block of the design is "Transformation" which cannot be obtained by any of the previously described methods. Therefore, the authors dedicated a first experiment to obtain the relationship between the LN  $\sigma$  and the CA-CFAR  $T$ , for a CA-CFAR having 64 cells in the reference window.

The methodology followed was similar to that described in [20, 21]. An amount of 40  $\sigma$  values, uniformly distributed in the range between 0,025 and 1,25, were chosen and a group of one million LN samples was generated for each  $\sigma$  value. Figure 2 shows the structure of the created set of samples, which will be referred from now on as *Set A*.



**Fig. 2.** Structure of the Set A that contains 40 million Log-Normal Samples.

**Source:** The authors.

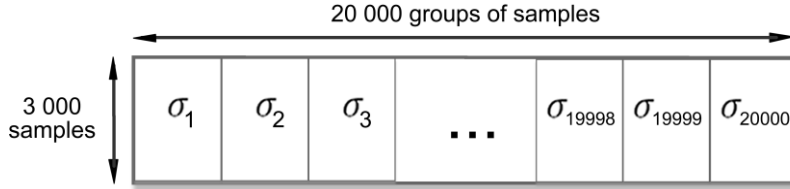
After generating the data, the first group from *Set A* was processed by a CA-CFAR operating with a low  $T$  value (for example  $T = 0,90$ ). Then, the  $P_f$  was measured. As the initial  $T$  value was low, the resulting  $P_f$  was quite high (for example 0,1). At this point, one iteration was completed.

In the next iteration,  $T$  was raised thereby producing a lower  $P_f$ . Many others iterations were executed until the exact values of  $T$  for were found for  $P_f = 10^{-2}$ ,  $P_f = 10^{-3}$  and  $P_f = 10^{-4}$  with a 1% deviation. At this point, the process ended for the first group of samples.

A similar experiment was performed on the remaining 39 groups, yielding a total of 120  $T$  values as a result of the test. With these values, a curve fitting procedure was carried out obtaining three mathematical expressions through which the optimal  $T$ s can be easily obtained for any  $\sigma$  and any of the three addressed  $P_f$ s. These three expressions were placed in the "Transformation" block from Figure 1.

### 1.3.2. MISTAKES COMMITTED BY THE METHOD OF MOMENTS

It was necessary to quantify the error committed by the "Estimation by MoM" block from Figure 1 in order to derive the overall accuracy of the scheme. For this purpose, the authors prepared a *Set B* composed of 20 000 groups of 3 000 samples each; see Figure 3. As it was done for *Set A*, the  $\sigma$  of each group was increased uniformly in the established interval.



**Fig. 3.** Structure of Set B that contains 60 million Log-Normal Samples.

**Source:** The authors.

Each group from *Set B* was processed with the MoM according to expressions (4) and (5). The incurred error was quantified and arranged into histograms for its better characterization.

### 1.3.3. INFLUENCE OF THE MOM ERROR OVER THE SELECTION OF T

The influence of the MoM error in the CA-CFAR multiplier ( $T$ ) selection was studied next. *Set B* itself was used in the trials.

The test was divided in two parts. In the first, samples were evaluated with a CA-CFAR with a priori knowledge of the  $\sigma$  of each group. In the second, the procedure was repeated using the MoM  $\sigma$  estimation instead of the original  $\sigma$ . The difference between the two obtained  $T$  revealed the committed error.

### 1.3.4. INFLUENCE OF T SELECTION ON THE FALSE ALARM RATE

After knowing the mistake made in  $T$  selection, a final experiment was conducted to find the deviation of the operational  $P_f$  from the design value. As it was expected, a higher error in the selection of  $T$  resulted in an increased deviation of the  $P_f$ .

However, the error does not spread with an easy to deduce mechanism as the decision taken by the detector is always drastic. Regardless of the margin by which the decision was made, the result remains the same as the detector always choose between 1 and 0, between target and clutter. There are no intermediate levels.

For this test, a variation of *Set B*, named *Set C*, was used. *Set C* originates from separating the 20 000 groups from *Set A* into 20 sections of 1000 groups each. The purpose of these sections is to lighten the processing load and they didn't introduce any changes in the data.

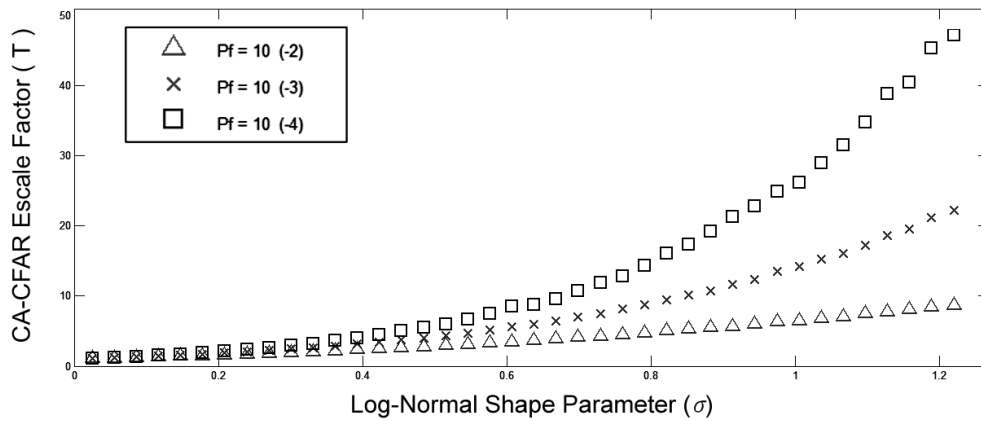
Each section from Set C was processed with a CA-CFAR whose  $T$  was defined according to estimates made by the MoM for each of the considered  $P_f$ s. The difference between the expected and the achieved  $P_f$  revealed the deviation introduced by the MoM and the global precision with which operates the LN-MoM-CA-CFAR solution.

## 2. RESULTS AND DISCUSSION

The "Results and Discussion" section will first show how the "Transformation" block from Figure 1 was constructed. The construction of the block is, in fact, the most important contribution of the research as it was the main obstacle in the creation of the LN-MoM-CA-CFAR detector. Secondly, the section will provide evidence of the proper functioning of the new scheme, describing the error introduced by its individual components and the reduced deviation that it achieves for the operational  $P_f$ .

### 2.1. INFLUENCE OF $T$ SELECTION ON THE FALSE ALARM RATE

By performing Monte Carlo simulations with the 40 millions of samples contained in Set A, a total of 120  $T$  values were obtained for the three addressed  $P_f$ s. Figure 4 plots the  $T$  values, revealing the remarkable influence of  $\sigma$  over  $T$  in the  $\sigma > 0,8$  region.



**Fig. 4.** Optimal  $T$  values for multiple statistical conditions of Log-Normal distributed clutter.

**Source:** The authors.

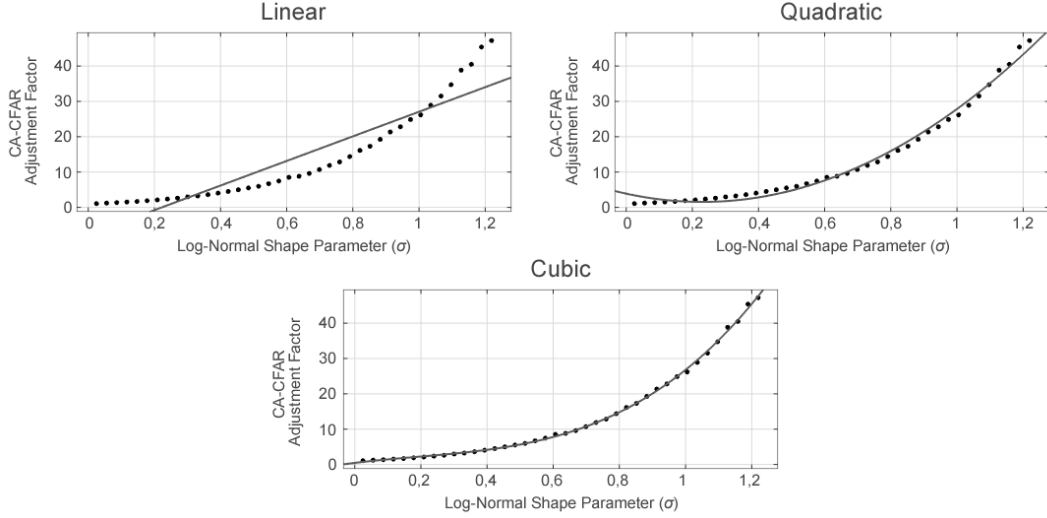
The cause of the observed tendency lies in the appearance of heavy tails for the higher figure of  $\sigma$ . A heavy tail produces sporadic high magnitude samples that disturb the most the lower  $P_f$ s.

In order to build the "Transformation" block, mathematical expressions that provide  $T$  values for any occurrence of  $\sigma$  in the 0,025 – 1,25 region are required. These expressions will avoid the storage of the values plotted in Figure 4 and will also provide a soft approximation to any  $\sigma$  within the interval.

After trying different alternatives, the authors concluded that the polynomial fits displayed the best resemblance with the data. Figure 5 shows, the polynomials



fits of first, second and third order for the  $P_f = 10^{-4}$  case. The linear approximation was obviously inappropriate since it exhibited no resemblance with the data. The quadratic fit was a little closer, but there were several areas where the departure was significant. Instead, the cubic polynomial followed very closely the behavior of the data.



**Fig. 5.** Low order polynomial fits for data from Figure 4 corresponding to  $P_f = 10^{-4}$ .

**Source:** The authors.

In short, fits from second to eighth order were considered, being the incurred mistake made showed in Table 1 for  $P_f = 10^{-4}$ . Note that a constant improvement in accuracy is achieved up until the fourth order, where the RMSE (Root Mean Square Error) and SSE (Sum of Squared Error) reach a saturation point. A similar behavior was observed for data corresponding to  $P_f = 10^{-2}$  and  $P_f = 10^{-3}$ .

**Table 1:** Results of the polynomial fit for  $P_f = 10^{-4}$ .

Polynomial Degree	RMS	SSE
2	1,286	6,12
3	0,4058	5,928
4	0,3872	5,055
5	0,3852	5,046
6	0,3896	5,009
7	0,3869	4,791
8	0,3781	4,576

**Source:** The authors.

Consequently, the authors selected the fourth-order polynomial fit for the expressions that composed the "Transformation" block from Figure 1.

Therefore, the CA-CFAR optimal  $T$ s can be obtained through (6), (7) and (8), corresponding to the  $P_f$ s of  $10^{-4}$ ,  $10^{-3}$  and  $10^{-2}$  respectively

$$T = 13,91\sigma^4 - 0,624\sigma^3 + 8,729\sigma^2 + 3,608\sigma + 0,994 \quad (6)$$

$$T = 2,56\sigma^4 + 2,407\sigma^3 + 5,247\sigma^2 + 2,839\sigma + 1,018 \quad (7)$$

$$T = -0,1728\sigma^4 + 0,8586\sigma^3 + 2,558\sigma^2 + 2,236\sigma + 1,007 \quad (8)$$

Alternatively, it was found that the fit for rational expressions, such as (9), also achieved a good approximation with the data.

$$f(x) = \frac{(p_1x + p_2)}{(x^2 + q_1x + q_2)} \quad (9)$$

However, the RMSE of the rational fit was slightly higher compared to the fourth order polynomial RMSE. Thus, the latter was preferred.

## 2.2. ERROR IN THE LN SHAPE PARAMETER ESTIMATION

When the  $\sigma$  of the clutter is known a priori, the optimal  $T$  can be chosen to operate with the desired  $P_f$  using expressions (6), (7) and (8). However, in a real operating environment,  $\sigma$  is unknown and the proposed scheme performs its estimation through the MoM. The resulting estimate will be slightly deviated from the real value, because a finite set of samples was used in the process.

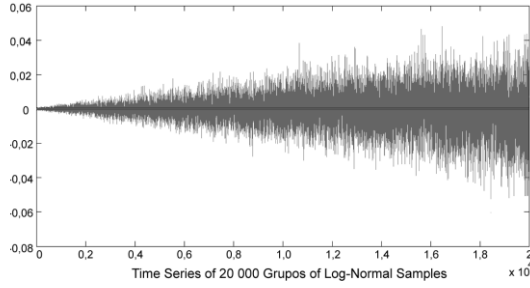
Figure 6 shows the error committed by the LN-MoM-CA-CFAR in the estimation of  $\sigma$  after evaluating Set A for  $P_f = 10^{-4}$ . The results are appropriate, since they are equally distributed in positive and negative magnitudes. This is easily visible in Figure 7 that plots a histogram of the errors.

The average error was 0,006566 and the maximum error was 0,048. These are excellent results because they represent mistakes of only 0,5% and 5% respectively.

## 2.3. DEVIATION OF THE ADJUSTMENT FACTOR

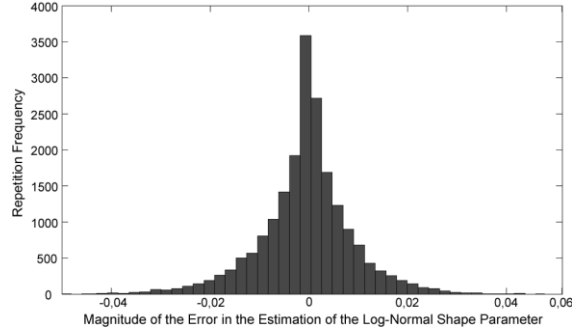
The error incurred in the estimation of  $\sigma$  affects the choice of the adjustment factor ( $T$ ). In order to determine the deviation of the  $T$  estimates, experiments were conducted with Set B. Figure 8 shows the deviation of the absolute error in the adjustment factor for  $P_f = 10^{-4}$ . Figure 9 does the same with the percentage of the relative error.

The mean absolute error in the selection of  $T$  for  $P_f = 10^{-4}$  was 0,3946 and the maximum deviation was 6,6113. The average relative error was 1,9409% and the maximum relative error was 14,8785%. For  $P_f = 10^{-3}$  the figures were 0,1655, 2,4196, 1,5368% and 11,8329%, and for  $P_f = 10^{-2}$ : 0,0523, 0,6190, 1,0345% and 7,7258%. Summarizing, it's safe to state that the average mistakes are less than a 2% and the maximum mistakes less than a 15%.



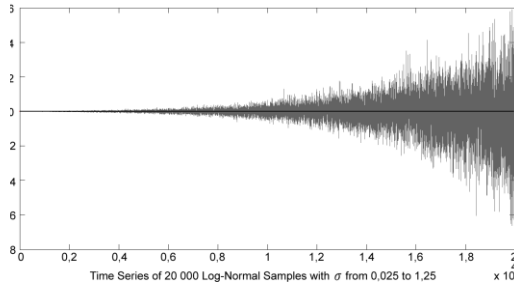
**Fig. 6.** Magnitude of the errors committed in the estimation of the Log-Normal shape parameter ( $\sigma$ ).

**Source:** The authors.



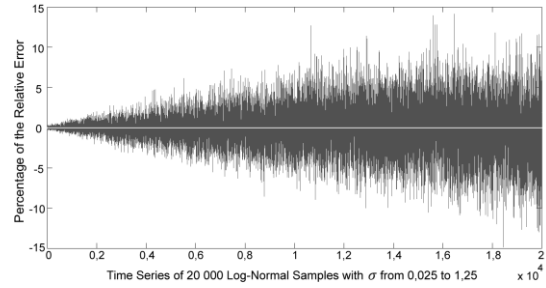
**Fig. 7.** Histogram of the error in the estimation of the Log-Normal shape parameter ( $\sigma$ ).

**Source:** The authors.



**Fig. 8.** Absolute error in the selection of the CA-CFAR adjustment factor ( $T$ ).

**Source:** The authors.



**Fig. 9.** Relative error in the selection of the CA-CFAR adjustment factor ( $T$ ).

**Source:** The authors.

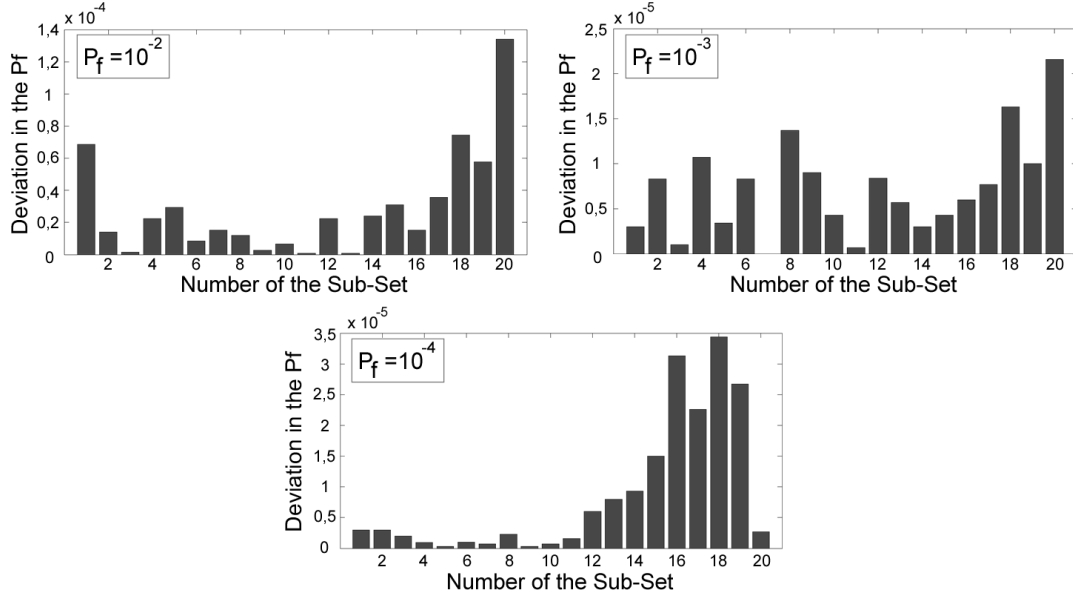
## 2.4. DEVIATION IN THE FALSE ALARM PROBABILITY

The objective of the new LN-MoM-CA-CFAR is to operate with a minimum deviation of the  $P_f$  from the design value. The previous tests characterized the deviation of the internal components of the solution. The last test, whose results are shown ahead, used the internal partial results to estimate the overall deviation of the  $P_f$ .

Figure 10 summarizes the experiments conducted with *Set C*. Results are plotted for  $P_f = 10^{-2}$ ,  $P_f = 10^{-3}$  and  $P_f = 10^{-4}$ . Each bin describes the  $P_f$  deviation experienced after processing 1000 groups of 3000 samples each. The bins to the left in each figure represent the lowest values of  $\sigma$ . Consequently, the bins to the right display a higher magnitude because they represent situations where the detector deals with clutter having a very heavy tailed distribution.

For  $P_f = 10^{-4}$ , the average deviation of the operational false alarm probability was  $8,5935 \cdot 10^{-6}$ , which represented an 8,5935% deviation from the design  $P_f$ . For  $P_f = 10^{-3}$ , the average departure was  $7,27 \cdot 10^{-6}$ , for a 0,7270%. Lastly, the

average deviation was  $2,884 \cdot 10^{-5}$  for  $P_f = 10^{-2}$ , which represents a 0,2884% deviation.

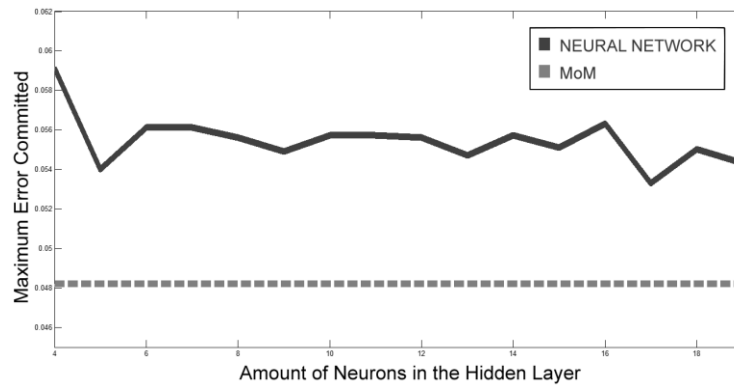


**Fig. 10.** Deviation of the False Alarm Probability of sub-sets containing 1 000 groups of 3000 samples.

**Source:** The authors.

## 2.5. COMPARING RESULTS

The deviation obtained by the LN-MoM-CA-CFAR processor from the design  $P_f$  is very low. This can be appreciated if the results are compared with a similar implementation. In [34] the accuracy of the MoM for the Weibull distribution was measured and an improved scheme was proposed using neural networks. The average deviations achieved were 4,56%, 14,92% and 34,52% for the  $P_{fs}$  of  $10^{-2}$ ,  $10^{-3}$  y  $10^{-4}$  respectively. Note that these values are much worse than those achieved for the Log-Normal case (0,2884%, 0,7270% and 8,5935%).



**Fig. 11.** Comparison between the maximum error of the neural and MoM alternatives.

**Source:** The authors.

Additionally, a neuronal solution similar to that proposed in [34] was developed by the authors, but no satisfactory results were found. Figure 11 shows how the

variation of the number of neurons in the hidden layer, main design criteria used in [34], fails to provoke an error inferior to the one exhibited by the MoM in the estimation of  $\sigma$ .

Additional simulations were conducted with a CA-CFAR with no correction of the scale factor, and it was observed that the processor experienced deviations superior to 200% even when the whole interval of possible values of  $\sigma$  was not swept. The authors could not find any other similar system in the related literature dealing with the problem of correcting the adjustment factor for guaranteeing a reduced deviation from the design false alarm probability. So, they consider that the current proposal is innovative in its field since it solves a problem often ignored by new CFAR schemes, which are usually concentrated of modifying the method for estimating the average of the background.

### 3. CONCLUSIONS AND FUTURE RESEARCH

The new LN-MoM-CA-CFAR radar detector, capable of maintaining a reduced deviation from the design false alarm probability even when facing statistically fluctuating Log-Normal clutter, was presented. The solution solves a problem of proven influence on the performance of radar processors and often ignored in most CFAR proposals.

The stability of the system is excellent. It's able to operate with a deviation of only 0,2884% for  $P_f = 10^{-2}$ . The results were achieved after performing a curve fitting procedure with data obtained by processing 100 million Log-Normal computer generated samples.

The authors will focus next on developing similar solutions for the popular Compound Gaussian and Pareto statistical distributions that have been related to radar investigations for land and sea clutter. It's also interesting the adaptation of the scheme to the OS-CFAR (Ordered Statistics CFAR) alternative, which would begin to develop a methodology to design CFAR detectors adapted to changes in the clutter.

### 4. ACKNOWLEDGMENT

The authors thank the help offered by PhD Nelson Chávez Ferry regarding the theoretical radar fundamentals and the general conception of the project.

### 5. REFERENCES

- [1] Richards, M. A., Scheer, J. A. and Holm, W. A. (2010), *Principles of Modern Radar Vol I Basic Principles*. New Jersey, United States of America: Scitech Publishing.
- [2] Melvin, W. L. and Scheer, J. A., *Principles of Modern Radar (2014), Vol III Radar Applications*. New Jersey, United States of America: Scitech Publishing.
- [3] Ward, K., Tough, R. and Watts, S. (2013), *Sea Clutter Scattering, the K Distribution and Radar Performance*, (2nd edition). London, United Kingdom: The Institution of Engineering and Technology. doi: 10.1049/pbra025e.

- [4] Skolnik, M. I. (2008), *Radar Handbook* (3er Edition). New York, United States: McGraw-Hill, 2008.
- [5] Meikle, H. (2008), *Modern Radar Systems* (2nd Edition). Boston, United States: Artech House.
- [6] Rohling, H. (1983), Radar CFAR Thresholding in Clutter and Multiple Target Situations. *IEEE Transactions on Aerospace and Electronic Systems*, 19(4), pp.608-621. doi: 10.1109/TAES.1983.309350.
- [7] Qin, Y. and Gong, H. (2013), A New CFAR Detector based on Automatic Censoring Cell Averaging and Cell Averaging. *TELKOMNIKA*, 11(6), pp. 3298 - 3303.
- [8] Kumar Yadav, A. and Kant, L. (2013), Moving Target Detection using VI-CFAR Algorithm on MATLAB Platform. *International Journal of Advanced Research in Computer Science and Software Engineering*, 3(12), pp. 915-918.
- [9] El Mashade, M. B. (2013), Performance Analysis of the Modified Versions of CFAR Detectors in Multiple-Target and Nonuniform Clutter. *Radioelectronics and Communications Systems*, 56(8), pp. 385-401. doi: 10.3103/S0735272713080013
- [10] Shin, J. W. and Seo, Y. K. (2012), "Modified Variability-Index CFAR Detection Robust to Heterogeneous Environment". Presented in *the International Conference on Systems and Electronic Engineering*, Phuket, Thailand, Dec. 18-19, pp. 6-12.
- [11] Yim, J. Z., Chou, C. R. and Wong, W. K. (2007), "A Study of the Statistics of Sea Clutter in the Northern Coast of Taiwan". Presented in *Proceedings of the Seventeenth International Offshore and Polar Engineering Conference*, Lisbon, Portugal, pp. 1-6.
- [12] Ping, Q. (2011), "Analysis of Ocean Clutter for Wide-Band Radar Based on Real Data". Presented in *Proceedings of the 2011 International Conference on Innovative Computing and Cloud Computing*, Wuhan, China, pp. 121-124.
- [13] Chen, Z., Liu, X. and Wu, Z. (2013), "The Analysis of Sea Clutter Statistics Characteristics Based on the Observed Sea Clutter of Ku-Band Radar". Presented in *IEEE Proceedings of the International Symposium on Antennas & Propagation*, pp. 1183-1186.
- [14] Mezache, A., Soltani, F. and Sahed, M. (2013), "A Model for Non Rayleigh Sea Clutter Amplitudes using Compound Inverse Gaussian". Presented in *IEEE Radar Conference (RadarCon13)*, Apr. 29-May 3. doi: 10.1109/RADAR.2013.6585989
- [15] Machado Fernández, J. R. and Bacallao Vidal, J. C. (2014), MATE-CFAR: Ambiente de Pruebas para Detectores CFAR en MATLAB. *Telem@tica*, 13(3), pp. 86-98.
- [16] Sayama, S. and Ishii, S. (2013), Suppression of Log-Normal Distributed Weather Clutter Observed by an S-Band Radar. *Wireless Engineering and Technology*, 4(3), pp. 125-133. doi: 10.4236/wet.2013.43019
- [17] Dong, Y. (2006), Distribution of X-Band High Resolution and High Grazing Angle Sea Clutter, *Technical Report DSTO-RR-0316*, Electronic Warfare and Radar Division, Defence Science and Technology Organization, Edinburgh, South Australia.
- [18] Machado Fernández, J. R. and Bacallao Vidal, J. C. (2016), Improved Shape Parameter Estimation in K Clutter with Neural Networks and Deep Learning. *International Journal of Interactive Multimedia and Artificial Intelligence*, 3(7), pp. 96-103. doi: 10.9781/ijimai.2016.3714
- [19] McLeod, J. W. (1998), *An Investigation of the CDF-Based Method of Moments*, Thesis for the degree of Master in Applied Science, Department of Civil Engineering, University of Toronto.
- [20] Machado Fernández, J. R. (2015), Estimation of the Relation between Weibull Distributed Sea clutter and the CA-CFAR Scale Factor. *Journal of Tropical Engineering*, 25(2), pp. 19-28. doi: 10.15517/jte.v25i2.18209

- [21] Machado Fernández, J. R. and Bacallao Vidal, J. C. (2016), Optimal Selection of the CA-CFAR Adjustment Factor for K Power Sea Clutter with Statistical Variations, (accepted) *Ciencia e Ingeniería Neogranadina*.
- [22] Sayama, S. and Ishii, S. (2011), Amplitude Statistics of Sea Clutter by MDL Principle. *IEEJ Transactions Fundamentals and Materials*, 132(10), pp. 886-892. doi: 10.1541/ieejfms.132.886
- [23] Rajalakshmi Menon, K., Balakrishnan, N. and Janakiraman, M. (1995), Characterization of Fluctuation Statistics of Radar Clutter for Indian Terrain. *IEEE Transactions on Geoscience and Remote Sensing*, 33(2), pp. 317 - 324. doi: 10.1109/36.377931
- [24] Sayama, S. and Ishii, S. (2011), Amplitude Statistics of Ground Clutter from Town and Hill Observed by S-band Radar, *IEEJ Transactions Fundamentals and Materials*, 131(11), pp. 916-923. doi: 10.1541/ieejfms.131.916
- [25] Dong, Y. (2004), Clutter Spatial Distribution and New Approaches of Parameter Estimation for Weibull and K-Distributions, *DSTO-RR-0274*, DSTO Systems Sciences Laboratory, Edingburgh, South Australia, Australia.
- [26] Stehwien, W. (1994), Statistical and Correlation Properties of High Resolution X-band Sea Clutter. Presented in '94 *IEEE Proceedings*, pp. 46-51, 1994. doi: 10.1109/NRC.1994.328096
- [27] Farina, A., Gini, F. and Greco, M. V. (1997), High Resolution Sea Clutter Data: Statistical Analysis of Recorded Live Data. *IEE Proceedings on Radar, Sonar and Navigation*, 144(3), pp. 121-130. doi: 10.1049/ip-rsn:19971107
- [28] Ishii, S., Sayama, S. and Mizutani, K. (2011), Effect of Changes in Sea-Surface State on Statistical Characteristics of Sea Clutter with X-band Radar. *Wireless Engineering and Technology*, 2(3), pp. 175-183. doi: 10.4236/wet.2011.23025
- [29] Weiping, N., Weidong, Y. and Hui, B. (2013), Statistical Analysis of High Resolution TerraSAR-X Images for Ground Targets Detection. Presented in *2013 Seventh International Conference on Image and Graphics*, pp. 343-347, Qingdao, China, Jul. 26-28. doi: 10.1109/ICIG.2013.74
- [30] Jurgens Strydom, J. (2012), *Generic Ground Clutter Simulation for Radar Testing and Evaluation*, Master Thesis, Radar Remote Sensing Group, Department of Electrical Engineering, University of Cape Town, 2012.
- [31] Ward, K., Tough, R. and Watts, S. (2006), *Sea Clutter. Scattering, the K distribution and Radar Performance*, London, United Kingdom: The Institution of Engineering and Technology.
- [32] Greco, M., Bordonì, F. and Gini, F. (2004), X-Band Sea-Clutter nonstationarity: Influence of Long Waves, *IEEE Journal of Oceanic Engineering*, 29(2), pp. 269-283. doi: 10.1109/JOE.2004.828548
- [33] Sayama, S. and Sekine, M. (2001), Weibull, Log-Weibull and K-Distributed Ground Clutter Modelling Analyzed by AIC, *IEEE Transactions on Aerospace and Electronic Systems*, 37(3), pp. 1108-1113. doi: 10.1109/7.953262
- [34] Machado Fernández, J. R. and Bacallao Vidal, J. C. (2016), Procesador de Promediación con Corrección Adaptativa del Factor de Ajuste Mediante Redes Neuronales Artificiales, Presented in *VIII Congreso Internacional de Telemática y Telecomunicaciones, Convención Científica de Ingeniería y Arquitectura '16, Palacio de las Convenciones*, Oct. 21-25, La Habana, Cuba.